

A Class of Small Deviation Theorems for Random Sums of Multivariate Function of m th-Order Markov Chain

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Abstract. In this paper, the notion of limit relative logarithmic likelihood ratio of stochastic sequence, as a measure of dissimilarity between the joint distribution and the m th-order nonhomogeneous Markov distribution, is introduced. A kind of strong limit theorems represented by inequalities which we call the small deviation theorems for the arbitrary stochastic sequence are obtained by constructing the consistent distribution functions. As corollaries, we obtain some small deviation theorems for the occurrence frequency of the state groups and the harmonic mean of the transitional probabilities of the m th-order nonhomogeneous Markov chain.

Key Words and Phrases: stochastic sequence; limit logarithmic likelihood ratio; small deviation theorem; m th-order nonhomogeneous Markov chain; random sum

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1. Introduction

Let $\{X_n, n \geq 0\}$ be an arbitrary stochastic sequence defined on the probability space (Ω, \mathcal{F}, P) which takes values in an alphabet set $S = \{s_0, s_1, \dots, s_{N-1}\}$, with the joint distribution:

$$P(X_0 = x_0, \dots, X_n = x_n) = p(x_0, \dots, x_n) > 0, \quad x_i \in S, \quad 0 \leq i \leq n. \quad (1)$$

We obtain by the definition of the conditional probability

$$p(x_0, \dots, x_n) = p(x_0) \prod_{k=1}^n p_k(x_k | x_0, \dots, x_{k-1}). \quad (2)$$

Let Q be another probability measure on (Ω, \mathcal{F}) , $\{X_n, n \geq 0\}$ be an m th-order nonhomogeneous Markov chain on the measure Q , with the m -dimensional initial distribution and m th-order transition probabilities as follows:

$$q_0(i_0, \dots, i_{m-1}) = Q(X_0 = i_0, \dots, X_{m-1} = i_{m-1}), \quad i_0, \dots, i_{m-1} \in S, \quad (3)$$

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$$q_n(j|i_1, \dots, i_m) = Q_n(X_n = j | X_{n-m} = i_1, \dots, X_{n-1} = i_m), \quad i_1, \dots, i_m, j \in S, \quad n \geq m. \quad (4)$$

The joint distribution of $\{X_n, n \geq 0\}$ with respect to the measure Q is

$$Q(X_0 = x_0, \dots, X_n = x_n) = q(x_0, \dots, x_n) = q_o(x_0, \dots, x_{m-1}) \prod_{k=m}^n q_k(x_k | x_{k-m}, \dots, x_{k-1}),$$

$$x_i \in S, \quad 0 \leq i \leq n. \quad (5)$$

Definition 1. Let p and q be defined as (1) and (5), $\{\sigma_n, n \geq 0\}$ be a nonnegative increasing stochastic sequence and $\sigma_n \uparrow \infty$. Denote

$$h(P|Q) = \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \ln \left[\frac{p(X_0, \dots, X_{[\sigma_n]})}{q_o(X_0, \dots, X_{m-1}) \prod_{k=m}^{[\sigma_n]} q_k(X_k | X_{k-m}, \dots, X_{k-1})} \right], \quad (6)$$

where $[\sigma_n]$ represents the integral part of σ_n . $h(P|Q)$ is called the sample divergence rate of the measure P relative to the measure Q .

In fact, $h(P|Q)$ is also called the limit relative logarithmic likelihood ratio or asymptotic logarithmic likelihood ratio of $\{X_n, n \geq 0\}$ on the measure P relative to Q . Although $h(P|Q)$ is not an optimum metric between the two probability measures, we nevertheless think of it as a measure of "dissimilarity" between their joint distribution and the m th-order nonhomogeneous Markov distribution. Obviously, $h(P|Q) = 0$ if and only if $P = Q$. It will be shown in (20) that $h(P|Q) \geq 0$, *a.s.(almostsurely)* in any case. Hence, $h(P|Q)$ can be used as a random measure of the deviation between the true joint distribution $p(x_0, \dots, x_n), (n \geq 0)$ and the m th-order nonhomogeneous Markov reference distribution $q_o(x_0, \dots, x_{m-1}) \prod_{k=m}^n q_k(x_k | x_{k-m}, \dots, x_{k-1})$. Roughly speaking, this deviation may be regarded as the one between $\{X_n, n \geq 0\}$ under the the measure P and the one under Q . The smaller $h(P|Q)$ is, the smaller the deviation is.

Study for strong limit properties of nonhomogeneous Markov chain is always one of central parts of the limit theory of probability theory. Many scholars have studied the subject until now. Liu and and Yang (see [1]) have studied the asymptotic equipartition properties (AEP) and limit properties of function sequence of nonhomogeneous Markov chain. Liu (see [2]) has discussed the strong limit theorems relative to the geometric average of random transition properties of finite nonhomogeneous Markov chain. Liu and Yang (see [3]) have investigated the strong deviation theorems of nonhomogeneous Markov chain relative to arbitrary stochastic sequence and AEP approximation of nonhomogeneous Markov information source. Yang (see [4]) has furthermore studied a class of small deviation theorems for the sequence of N -valued random variables with respect to m th-order nonhomogeneous Markov chains. Liu (see [5]) has discussed the strong limit theorems for the harmonic mean of random transition probabilities of nonhomogeneous Markov chain. Liu and Wang (see [6]) have proved the strong limit properties for the state couples of nonhomogeneous Markov chain on the random selection system. Wang (see [7]) has studied the AEP and limit theorems for nonhomogeneous Markov chain on the generalized gambling system. Afterward, many scholars (see [11-28]) have studied all kinds of

stochastic processes and some limit properties with their applications for nonhomogeneous Markov chains on the generalized gambling system.

Many of practical information sources, such as language and image information, are often m th-order Markov chain, and always nonhomogeneous. m th-order nonhomogeneous Markov chain is a natural generalization of the general nonhomogeneous Markov chain. Hence it is of importance to study the limit properties for the m th-order nonhomogeneous Markov chain in the information theory and the probability theory. Yang and Liu (see [9]) have proved the limit theorem for averages of the functions of $m + 1$ variables of m th-order nonhomogeneous Markov chain and the AEP for m th-order nonhomogeneous Markov information source. Wang (see [10]) has discussed the Shannon-McMillan theorems for m th-order nonhomogeneous Markov information source.

In this paper, our aim is to establish a class of small deviation theorems represented by inequalities for the dependent stochastic sequence with respect to m th-order nonhomogeneous Markov chain by introducing the sample relative entropy rate as a measure of deviation between the arbitrary stochastic sequence and the m th-order nonhomogeneous Markov chain. We apply a new type of techniques distinct from that of Liu and Yang (see [4] and [9]) to study the small deviation theorems for arbitrary stochastic sequence. As corollaries, a class of small deviation theorems for the occurrence frequencies of state groups and some limit properties for harmonic mean of the transitional probabilities of the m th-order Markov chain are obtained.

We denote $X_m^n = \{X_m, \dots, X_n\}$, $i_0^m = \{i_0, \dots, i_m\}$. x_m^n is the realization of X_m^n .

2. Main Result and Its Proof

Theorem 1. *Let $\{X_n, n \geq 0\}$ be an arbitrary stochastic sequence with the joint distribution (1), $h(P|Q)$ be defined by (6), $\{\sigma_n, n \geq 0\}$ be an increasing nonnegative stochastic sequence. Let $g(x_0, \dots, x_m)$ be a multivariate real function defined on S^{m+1} . Denote*

$$D(c) = \{\omega : \lim_{n \rightarrow \infty} \sigma_n(\omega) = +\infty, h(P|Q) \leq c\}. \quad (7)$$

Then

$$\limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \{g(X_{k-m}^k) - E_Q(g(X_{k-m}^k)|X_{k-m}^{k-1})\} \leq (2\sqrt{c} + c) \sum_{i_0 \cdots i_m \in S} |g(i_0, \dots, i_m)|, \quad (8)$$

$P - a.s. \quad \omega \in D(c)$

$$\liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \{g(X_{k-m}^k) - E_Q(g(X_{k-m}^k)|X_{k-m}^{k-1})\} \geq -2\sqrt{c} \sum_{i_0 \cdots i_m \in S} |g(i_0, \dots, i_m)|, \quad (9)$$

$P - a.s. \quad \omega \in D(c)$

where $[\sigma_n]$ represents the integral part of σ_n , and E_Q is the expectation with respect to the measure Q .

Proof. We consider the probability space (Ω, \mathcal{F}, P) . Let λ be an arbitrary real number, and $\delta_i(j)$ be the Kronecker function. Denote

$$\begin{aligned} \delta_{i_0 \cdots i_m}(x_{k-m}^k) &= \delta_{i_0}(x_{k-m}) \cdots \delta_{i_m}(x_k), \\ \delta_{i_0 \cdots i_{m-1}}(x_{k-m}^{k-1}) &= \delta_{i_0}(x_{k-m}) \cdots \delta_{i_{m-1}}(x_{k-1}), \\ M_k(x_{k-m}, \cdots, x_k; \lambda) &= \left(\frac{1}{1 + (\lambda - 1)q_k(i_m | i_0^{m-1})} \right)^{\delta_{i_0 \cdots i_{m-1}}(x_{k-m}^{k-1})} \lambda^{\delta_{i_0 \cdots i_m}(x_{k-m}^k)} q_k(x_k | x_{k-m}^{k-1}), \\ & \qquad \qquad \qquad k \geq m. \end{aligned} \quad (10)$$

We construct the following product distribution:

$$\mu(x_0, \cdots, x_n; \lambda) = q_0(x_0, \cdots, x_{m-1}) \prod_{k=m}^n M_k(x_{k-m}, \cdots, x_k; \lambda). \quad (11)$$

By (10) and (11) we have

$$\begin{aligned} & \sum_{x_n \in S} \mu(x_0, \cdots, x_n; \lambda) = \\ &= \sum_{x_n \in S} q_0(x_0^{m-1}) \prod_{k=m}^n \left(\frac{1}{1 + (\lambda - 1)q_k(i_m | i_0^{m-1})} \right)^{\delta_{i_0 \cdots i_{m-1}}(x_{k-m}^{k-1})} \lambda^{\delta_{i_0 \cdots i_m}(x_{k-m}^k)} q_k(x_k | x_{k-m}^{k-1}) \\ &= \mu(x_0, \cdots, x_{n-1}; \lambda) \sum_{x_n \in S} \left(\frac{1}{1 + (\lambda - 1)q_n(i_m | i_0^{m-1})} \right)^{\delta_{i_0 \cdots i_{m-1}}(x_{n-m}^{n-1})} \lambda^{\delta_{i_0 \cdots i_m}(x_{n-m}^n)} q_n(x_n | x_{n-m}^{n-1}) \\ &= \mu(x_0, \cdots, x_{n-1}; \lambda) \frac{1}{[1 + (\lambda - 1)q_n(i_m | i_0^{m-1})]^{\delta_{i_0 \cdots i_{m-1}}(x_{n-m}^{n-1})}} \left(\sum_{x_n = i_m} + \sum_{x_n \neq i_m} \right) \\ &= \mu(x_0, \cdots, x_{n-1}; \lambda) \frac{\lambda^{\delta_{i_0 \cdots i_{m-1}}(x_{n-m}^{n-1})} q_n(i_m | x_{n-m}^{n-1}) + 1 - q_n(i_m | x_{n-m}^{n-1})}{[1 + (\lambda - 1)q_n(i_m | i_0^{m-1})]^{\delta_{i_0 \cdots i_{m-1}}(x_{n-m}^{n-1})}}. \end{aligned} \quad (12)$$

When $\delta_{i_0 \cdots i_{m-1}}(x_{n-m}^{n-1}) = 0$, we obtain from (12) that

$$\begin{aligned} & \sum_{x_n \in S} \mu(x_0, \cdots, x_n; \lambda) = \\ &= \mu(x_0, \cdots, x_{n-1}; \lambda) \frac{q_n(i_m | x_{n-m}^{n-1}) + 1 - q_n(i_m | x_{n-m}^{n-1})}{[1 + (\lambda - 1)q_n(i_m | i_0^{m-1})]^0} = \mu(x_0, \cdots, x_{n-1}; \lambda). \end{aligned} \quad (13)$$

When $\delta_{i_0 \dots i_{m-1}}(x_{n-m}^{n-1}) = 1$, we acquire from (12) that

$$\begin{aligned} & \sum_{x_n \in S} \mu(x_0, \dots, x_n; \lambda) \\ &= \mu(x_0, \dots, x_{n-1}; \lambda) \frac{1 + (\lambda - 1)q_n(i_m | i_0^{m-1})}{1 + (\lambda - 1)q_n(i_m | i_0^{m-1})} = \mu(x_0, \dots, x_{n-1}; \lambda). \end{aligned} \quad (14)$$

Therefore, $\mu(x_0, \dots, x_n; \lambda)$, $n = 1, 2, \dots$ are a family of consistent distribution functions on S^{n+1} . Denote

$$T_n(\lambda, \omega) = \frac{\mu(X_0, \dots, X_n; \lambda)}{p(X_0, \dots, X_n)}. \quad (15)$$

By (5), (11) and (15), we can rewrite (15) as

$$\begin{aligned} T_n(\lambda, \omega) &= \lambda^{\sum_{k=m}^n \delta_{i_0 \dots i_m}(X_{k-m}^k)} \prod_{k=m}^n \left[\frac{1}{1 + (\lambda - 1)q_k(i_m | i_0^{m-1})} \right]^{\delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1})} \\ &\quad \cdot q_o(X_0^{m-1}) \prod_{k=m}^n q_k(X_k | X_{k-m}^{k-1}) / p(X_0, \dots, X_n). \end{aligned} \quad (16)$$

Since μ and p are two probability measures, we can know that $T_n(\lambda, \omega)$ is a nonnegative sup-martingale from Doob's martingale convergence theorem (see [8]). Moreover,

$$\lim_{n \rightarrow \infty} T_n(\lambda, \omega) = T_\infty(\lambda, \omega) < \infty. \quad P - a.s. \quad (17)$$

By (7) and (17) we have

$$\limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \ln T_{[\sigma_n]}(\lambda, \omega) \leq 0. \quad P - a.s. \quad \omega \in D(c) \quad (18)$$

By (16) and (18) we can obtain

$$\begin{aligned} & \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \left\{ \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_m}(X_{k-m}^k) \ln \lambda - \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1}) \ln [1 + (\lambda - 1)q_k(i_m | i_0^{m-1})] \right. \\ & \left. - \ln \left[p(X_0, \dots, X_{[\sigma_n]}) / q_o(X_0^{m-1}) \prod_{k=m}^{[\sigma_n]} q_k(X_k | X_{k-m}^{k-1}) \right] \right\} \leq 0. P - a.s. \quad \omega \in D(c) \quad (19) \end{aligned}$$

Letting $\lambda = 1$ in (19), by (6) we can know

$$h(P|Q) \geq \liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \ln \left[\frac{p(X_0, \dots, X_{[\sigma_n]})}{q_o(X_0, \dots, X_{m-1}) \prod_{k=m}^{[\sigma_n]} q_k(X_k | X_{k-m}^{k-1})} \right] \geq 0. \quad P - a.s. \quad \omega \in D(c) \quad (20)$$

By (6), (7) and (19), taking into account the property of superior limit, we can write

$$\begin{aligned} \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \left\{ \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_m} (X_{k-m}^k) \ln \lambda - \right. \\ \left. \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_{m-1}} (X_{k-m}^{k-1}) \ln [1 + (\lambda - 1)q_k(i_m | i_0^{m-1})] \right\} \leq h(P|Q) \leq c. \end{aligned}$$

$$P - a.s. \quad \omega \in D(c) \quad . \quad (21)$$

In the case $\lambda > 1$, dividing two sides of (21) by $\ln \lambda$, we have

$$\limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \left\{ \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_m} (X_{k-m}^k) - \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_{m-1}} (X_{k-m}^{k-1}) \frac{\ln [1 + (\lambda - 1)q_k(i_m | i_0^{m-1})]}{\ln \lambda} \right\} \leq \frac{c}{\ln \lambda}.$$

$$P - a.s. \quad \omega \in D(c) \quad (22)$$

By (22), using the inequalities $1 - 1/x \leq \ln x \leq x - 1, (x > 0), 0 \leq \delta_{i_0, \dots, i_{m-1}}(X_{k-m}^{k-1}; \omega) \leq 1$ and the properties of superior limit

$$\limsup_{n \rightarrow \infty} (a_n - b_n) \leq d \Rightarrow \limsup_{n \rightarrow \infty} (a_n - c_n) \leq \limsup_{n \rightarrow \infty} (b_n - c_n) + d,$$

we can write

$$\begin{aligned} & \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} [\delta_{i_0 \dots i_m} (X_{k-m}^k) - \delta_{i_0 \dots i_{m-1}} (X_{k-m}^{k-1}) q_k(i_m | i_0^{m-1})] \\ & \leq \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \left\{ \delta_{i_0 \dots i_{m-1}} (X_{k-m}^{k-1}) \frac{\ln [1 + (\lambda - 1)q_k(i_m | i_0^{m-1})]}{\ln \lambda} \right. \\ & \quad \left. - \delta_{i_0 \dots i_{m-1}} (X_{k-m}^{k-1}) q_k(i_m | i_0^{m-1}) \right\} + \frac{c}{\ln \lambda} \\ & \leq \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \left\{ \delta_{i_0 \dots i_{m-1}} (X_{k-m}^{k-1}) \left[\frac{(\lambda - 1)q_k(i_m | i_0^{m-1})}{\ln \lambda} - q_k(i_m | i_0^{m-1}) \right] \right\} + \frac{c}{\ln \lambda} \\ & = \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_{m-1}} (X_{k-m}^{k-1}) q_k(i_m | i_0^{m-1}) \left(\frac{\lambda - 1}{\ln \lambda} - 1 \right) + \frac{c}{\ln \lambda} \\ & \leq \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_{m-1}} (X_{k-m}^{k-1}) q_k(i_m | i_0^{m-1}) \left(\frac{\lambda - 1}{1 - 1/\lambda} - 1 \right) + \frac{c}{1 - 1/\lambda} \\ & \leq \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_{m-1}} (X_{k-m}^{k-1}) (\lambda - 1) + c + \frac{c}{\lambda - 1} \end{aligned}$$

$$\leq \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} (\lambda - 1) + c + \frac{c}{\lambda - 1} = (\lambda - 1) + c + \frac{c}{\lambda - 1}.$$

$$P - a.s. \quad \omega \in D(c) \quad (23)$$

It is easy to show that in the case $c > 0$, the function $f(\lambda) = (\lambda - 1) + c + \frac{c}{\lambda - 1}$ ($\lambda > 1$) attains its smallest value $f(1 + \sqrt{c}) = 2\sqrt{c} + c$ at $\lambda = 1 + \sqrt{c}$. Hence, letting $\lambda = 1 + \sqrt{c}$ in (23), we have

$$\limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} [\delta_{i_0 \dots i_m}(X_{k-m}^k) - \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1})q_k(i_m|i_0^{m-1})] \leq 2\sqrt{c} + c.$$

$$P - a.s. \quad \omega \in D(c) \quad (24)$$

In the case $c = 0$, (24) also follows from (23) by choosing $\lambda_i \rightarrow 1 + (i \rightarrow \infty)$.

In the case $0 < \lambda < 1$, dividing two sides of (21) by $\ln \lambda$, we obtain

$$\begin{aligned} & \liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} [\delta_{i_0 \dots i_m}(X_{k-m}^k) - \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1})q_k(i_m|i_0^{m-1})] \\ & \geq \liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1}) \left(\frac{\ln[1 + (\lambda - 1)q_k(i_m|i_0^{m-1})]}{\ln \lambda} - q_k(i_m|i_0^{m-1}) \right) + \frac{c}{\ln \lambda} \\ & \geq \liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1})q_k(i_m|i_0^{m-1}) \left(\frac{\lambda - 1}{1 - 1/\lambda} - 1 \right) + \frac{c}{\lambda - 1} \\ & \geq \liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1})(\lambda - 1) + \frac{c}{\lambda - 1} \geq (\lambda - 1) + \frac{c}{\lambda - 1} \end{aligned}$$

$$P - a.s. \quad \omega \in D(c) \quad (25)$$

It is easy to show that in the case $c > 0$, the function $h(\lambda) = (\lambda - 1) + \frac{c}{\lambda - 1}$ ($0 < \lambda < 1$) attains its largest value $h(1 - \sqrt{c}) = -2\sqrt{c}$ at $\lambda = 1 - \sqrt{c}$. Hence, letting $\lambda = 1 - \sqrt{c}$ in (25), we obtain

$$\liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} [\delta_{i_0 \dots i_m}(X_{k-m}^k) - \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1})q_k(i_m|i_0^{m-1})] \geq -2\sqrt{c}.$$

$$P - a.s. \quad \omega \in D(c) \quad (26)$$

In the case $c = 0$, (26) also follows from (25) by choosing $\lambda_i \rightarrow 1 - (i \rightarrow \infty)$.

It follows from (24) and (26) that for any real function $g(i_0, \dots, i_m)$,

$$\limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} g(i_0, \dots, i_m) [\delta_{i_0 \dots i_m}(X_{k-m}^k) - \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1})q_k(i_m|i_0^{m-1})]$$

$$\leq (2\sqrt{c} + c)|g(i_0, \dots, i_m)|, \quad P - a.s. \quad \omega \in D(c) \quad (27)$$

$$\begin{aligned} & \liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} g(i_0, \dots, i_m) [\delta_{i_0 \dots i_m}(X_{k-m}^k) - \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1}) q_k(i_m | i_0^{m-1})] \\ & \geq -2\sqrt{c}|g(i_0, \dots, i_m)|. \quad P - a.s. \quad \omega \in D(c) \end{aligned} \quad (28)$$

Note that

$$\begin{aligned} & g(X_{k-m}, \dots, X_k) - E_Q(g(X_{k-m}, \dots, X_k) | X_{k-m}, \dots, X_{k-1}) \\ & = g(X_{k-m}, \dots, X_k) - \sum_{i_m \in S} g(X_{k-m}, \dots, X_{k-1}, i_m) q_k(i_m | X_{k-m}, \dots, X_{k-1}) \\ & = \sum_{i_0 \dots i_m \in S} \delta_{i_0 \dots i_m}(X_{k-m}, \dots, X_k) g(i_0, \dots, i_m) \\ & \quad - \sum_{i_0 \dots i_m \in S} \delta_{i_0 \dots i_{m-1}}(X_{k-m}, \dots, X_{k-1}) g(i_0, \dots, i_m) q_k(i_m | i_0, \dots, i_{m-1}) \\ & = \sum_{i_0 \dots i_m \in S} \delta_{i_0 \dots i_{m-1}}(X_{k-m}, \dots, X_{k-1}) g(i_0, \dots, i_m) [\delta_{i_m}(X_k) - q_k(i_m | i_0, \dots, i_{m-1})] \\ & = \sum_{i_0 \dots i_m \in S} g(i_0, \dots, i_m) [\delta_{i_0 \dots i_m}(X_{k-m}^k) - \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1}) q_k(i_m | i_0^{m-1})]. \end{aligned} \quad (29)$$

By virtue of the properties of superior limit and inferior limit, combing (27)-(29), we obtain

$$\begin{aligned} & \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \{g(X_{k-m}^k) - E_Q(g(X_{k-m}^k) | X_{k-m}^{k-1})\} \\ & = \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \sum_{i_0 \dots i_m \in S} g(i_0, \dots, i_m) [\delta_{i_0 \dots i_m}(X_{k-m}^k) - \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1}) q_k(i_m | i_0^{m-1})] \\ & \leq \sum_{i_0 \dots i_m \in S} \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} g(i_0, \dots, i_m) [\delta_{i_0 \dots i_m}(X_{k-m}^k) - \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1}) q_k(i_m | i_0^{m-1})] \\ & \leq (2\sqrt{c} + c) \sum_{i_0 \dots i_m \in S} |g(i_0, \dots, i_m)|, \end{aligned} \quad (30)$$

$$\begin{aligned} & \liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \{g(X_{k-m}^k) - E_Q(g(X_{k-m}^k) | X_{k-m}^{k-1})\} \\ & = \liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \sum_{i_0 \dots i_m \in S} g(i_0, \dots, i_m) [\delta_{i_0 \dots i_m}(X_{k-m}^k) - \delta_{i_0 \dots i_{m-1}}(X_{k-m}^{k-1}) q_k(i_m | i_0^{m-1})] \end{aligned}$$

$$\begin{aligned}
&\geq \sum_{i_0 \cdots i_m \in S} \liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} g(i_0, \cdots, i_m) [\delta_{i_0 \cdots i_m}(X_{k-m}^k) - \delta_{i_0 \cdots i_{m-1}}(X_{k-m}^{k-1}) q_k(i_m | i_0^{m-1})] \\
&\geq -2\sqrt{c} \sum_{i_0 \cdots i_m \in S} |g(i_0, \cdots, i_m)|.
\end{aligned} \tag{31}$$

(8), (9) follow from (30) and (31), respectively. The proof is finished. ◀

Corollary 1. *Let $\{X_n, n \geq 0\}$ be an m th-order nonhomogeneous Markov chain with the m dimensional initial distribution (3) and m th-order transitional probabilities (4). Let $g(x_0, \cdots, x_m)$ be a multivariate real function defined on S^{m+1} . Then*

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=m}^n \{g(X_{k-m}, \cdots, X_k) - E(g(X_{k-m}, \cdots, X_k) | X_{k-m}, \cdots, X_{k-1})\} = 0. \quad P\text{-a.s.} \tag{32}$$

Proof. Letting $\sigma_n = n$, $n \geq 1$, $P \equiv Q$ in Theorem 1, we obtain $h(P|Q) \equiv 0$, $E_Q(g(X_{k-m}^k) | X_{k-m}^{k-1}) = E(g(X_{k-m}^k) | X_{k-m}^{k-1})$ at the moment. This implies that $D(0) = \Omega$ when $c = 0$. Let $c = 0$. Then (32) follows from (8) and (9). ◀

3. Some small deviation theorems for the occurrence frequency of the state groups

In equation (4), if for all $n \geq m$,

$$\begin{aligned}
&Q_n(X_n = j | X_{n-m} = i_1, \cdots, X_{n-1} = i_m) \\
&= Q(X_n = j | X_{n-m} = i_1, \cdots, X_{n-1} = i_m), \quad \forall i_1, \cdots, i_m, j \in S
\end{aligned}$$

we call $\{X_n, n \geq 0\}$ an m th-order homogeneous Markov chain on the measure Q .

Corollary 2. *Let $\{X_n, n \geq 0\}$ be an arbitrary stochastic sequence with the joint distribution (1), $h(P|Q)$ be defined as before. Denote by $S_n(j_0, \cdots, j_m)$ the number of the state groups (j_0, \cdots, j_m) occurring in $(X_0, \cdots, X_m), (X_1, \cdots, X_{m+1}), \cdots, (X_{n-m}, \cdots, X_n)$, and by $S_n(j_0, \cdots, j_{m-1})$ the number of the state groups (j_0, \cdots, j_{m-1}) occurring in $(X_0, \cdots, X_{m-1}), (X_1, \cdots, X_m), \cdots, (X_{n-m}, \cdots, X_{n-1})$. That is,*

$$S_n(j_0, \cdots, j_m) = \sum_{k=m}^n \delta_{j_0 \cdots j_m}(X_{k-m}, \cdots, X_k), \tag{33}$$

$$S_n(j_0, \cdots, j_{m-1}) = \sum_{k=m}^n \delta_{j_0 \cdots j_{m-1}}(X_{k-m}, \cdots, X_{k-1}). \tag{34}$$

We set

$$L(c) = \{\omega : h(P|Q) \leq c\}. \tag{35}$$

Then

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \{S_n(j_0, \dots, j_m) - S_n(j_0, \dots, j_{m-1})q(j_m|j_0^{m-1})\} \leq (2\sqrt{c} + c),$$

$$P - a.s. \quad \omega \in L(c) \quad (36)$$

$$\liminf_{n \rightarrow \infty} \frac{1}{n} \{S_n(j_0, \dots, j_m) - S_n(j_0, \dots, j_{m-1})q(j_m|j_0^{m-1})\} \geq -2\sqrt{c}.$$

$$P - a.s. \quad \omega \in L(c) \quad (37)$$

Proof. Letting $\sigma_n = n$, $n \geq 0$, $g(X_{k-m}^k) = \delta_{j_0 \dots j_m}(X_{k-m}^k)$, $q_k(X_k|X_{k-m}^{k-1}) = q(X_k|X_{k-m}^{k-1})$, $k \geq m$ in Theorem 1, we can obtain $L(c) = D(c)$ and write

$$\begin{aligned} & \sum_{k=m}^n \{g(X_{k-m}^k) - E_Q(g(X_{k-m}^k)|X_{k-m}^{k-1})\} \\ = & \sum_{k=m}^n \{\delta_{j_0 \dots j_m}(X_{k-m}^k) - E_Q(\delta_{j_0 \dots j_m}(X_{k-m}^k)|X_{k-m}^{k-1})\} \\ = & \sum_{k=m}^n \{\delta_{j_0 \dots j_m}(X_{k-m}^k) - \sum_{x_k \in S} \delta_{j_0 \dots j_m}(X_{k-m}^{k-1}, x_k)q(x_k|X_{k-m}^{k-1})\} \\ = & \sum_{k=m}^n \{\delta_{j_0 \dots j_m}(X_{k-m}^k) - \delta_{j_0 \dots j_{m-1}}(X_{k-m}^{k-1})q(j_m|X_{k-m}^{k-1})\} \\ = & \sum_{k=m}^n \{\delta_{j_0 \dots j_m}(X_{k-m}^k) - \delta_{j_0 \dots j_{m-1}}(X_{k-m}^{k-1})q(j_m|j_0^{m-1})\} \\ = & S_n(j_0, \dots, j_m) - S_n(j_0, \dots, j_{m-1})q(j_m|j_0^{m-1}) \end{aligned} \quad (38)$$

and

$$\sum_{i_0 \dots i_m \in S} |g(i_0, \dots, i_m)| = \sum_{i_0 \dots i_m \in S} |\delta_{j_0 \dots j_m}(i_0, \dots, i_m)| = 1. \quad (39)$$

Taking into account (38) and (39), we obtain (36), (37) from (8), (9) immediately. ◀

Corollary 3. Let $\{X_n, n \geq 0\}$ be an arbitrary stochastic sequence with the joint distribution (1), $h(P|Q)$ be defined as before. Denote by $S_n(j_m)$ the number of the state groups j_m occurring in X_0, \dots, X_n . That is,

$$S_n(j_m) = \sum_{k=0}^n \delta_{j_m}(X_k). \quad (40)$$

Then

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \{S_n(j_m) - \sum_{j_0, \dots, j_{m-1} \in S} S_n(j_0, \dots, j_{m-1})q(j_m|j_0^{m-1})\} \leq (2\sqrt{c} + c)N^m,$$

$$P - a.s. \quad \omega \in L(c) \quad (41)$$

$$\liminf_{n \rightarrow \infty} \frac{1}{n} \left\{ S_n(j_m) - \sum_{j_0, \dots, j_{m-1} \in S} S_n(j_0, \dots, j_{m-1}) q(j_m | j_0^{m-1}) \right\} \geq -2\sqrt{c}N^m.$$

$$P - a.s. \quad \omega \in L(c) \quad (42)$$

Proof. Letting $\sigma_n = n, n \geq 0$, $g(X_{k-m}^k) = \delta_{j_m}(X_k)$, $q_k(X_k | X_{k-m}^{k-1}) = q(X_k | X_{k-m}^{k-1})$ in Theorem 1, we can write

$$\begin{aligned} & \sum_{k=m}^n \{g(X_{k-m}^k) - E_Q(g(X_{k-m}^k) | X_{k-m}^{k-1})\} \\ = & \sum_{k=m}^n \{\delta_{j_m}(X_k) - E_Q(\delta_{j_m}(X_k) | X_{k-m}^{k-1})\} \\ = & \sum_{k=m}^n \{\delta_{j_m}(X_k) - \sum_{x_k \in S} \delta_{j_m}(x_k) q(x_k | X_{k-m}^{k-1})\} \\ = & \sum_{k=m}^n \{\delta_{j_m}(X_k) - \sum_{j_0 \dots j_{m-1} \in S} \sum_{x_k \in S} \delta_{j_0 \dots j_{m-1}}(X_{k-m}^{k-1}) \delta_{j_m}(x_k) q(x_k | j_0^{m-1})\} \\ = & \sum_{k=m}^n \{\delta_{j_m}(X_k) - \sum_{j_0 \dots j_{m-1} \in S} \delta_{j_0 \dots j_{m-1}}(X_{k-m}^{k-1}) q(j_m | j_0^{m-1})\} \\ = & \sum_{k=m}^n \delta_{j_m}(X_k) - \sum_{j_0 \dots j_{m-1} \in S} q(j_m | j_0^{m-1}) \sum_{k=m}^n \delta_{j_0 \dots j_{m-1}}(X_{k-m}^{k-1}) \\ = & S_n(j_m) - \sum_{k=0}^{m-1} \delta_{j_m}(X_k) - \sum_{j_0 \dots j_{m-1} \in S} S_n(j_0, \dots, j_{m-1}) q(j_m | j_0^{m-1}) \end{aligned} \quad (43)$$

and

$$\sum_{i_0 \dots i_m \in S} |g(i_0, \dots, i_m)| = \sum_{i_0 \dots i_m \in S} |\delta_{j_m}(i_m)| = \sum_{i_0 \dots i_{m-1} \in S} 1 = N^m. \quad (44)$$

Combining (43), (44) with (8) and (9), we obtain (41), (42) immediately. ◀

We present a small deviation theorem for harmonic mean of the transitional probability of the m th-order Markov chain as follows:

Theorem 2. *Let $\{X_n, n \geq 0\}$ be an arbitrary stochastic sequence with the joint distribution (1), $h(P|Q)$ be defined by (6), $\{\sigma_n, n \geq 0\}$ be an increasing nonnegative stochastic sequence. Then*

$$\limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} q(X_k | X_{k-m}^{k-1})^{-1} \leq N + (2\sqrt{c} + c) \sum_{i_0 \dots i_m \in S} q(i_m | i_0^{m-1})^{-1},$$

$$P - a.s. \quad \omega \in D(c), \quad (45)$$

$$\liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} q(X_k | X_{k-m}^{k-1})^{-1} \geq N - 2\sqrt{c} \sum_{i_0 \cdots i_m \in S} q(i_m | i_0^{m-1})^{-1},$$

$$P - a.s. \quad \omega \in D(c). \quad (46)$$

Proof. Letting $g(X_{k-m}^k) = q(X_k | X_{k-m}^{k-1})^{-1}$ in Theorem 1, by (8) we can write

$$\begin{aligned} & \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \{g(X_{k-m}^k) - E_Q(g(X_{k-m}^k) | X_{k-m}^{k-1})\} \\ & \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \{q(X_k | X_{k-m}^{k-1})^{-1} - E_Q(q(X_k | X_{k-m}^{k-1})^{-1} | X_{k-m}^{k-1})\} \\ & \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \{q(X_k | X_{k-m}^{k-1})^{-1} - \sum_{x_k \in S} q(x_k | X_{k-m}^{k-1})^{-1} q(x_k | X_{k-m}^{k-1})\} \\ & \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} \{q(X_k | X_{k-m}^{k-1})^{-1} - N\} = \\ & = \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} q(X_k | X_{k-m}^{k-1})^{-1} - \limsup_{n \rightarrow \infty} \frac{\sigma_n - m + 1}{\sigma_n} N \\ & = \limsup_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} q(X_k | X_{k-m}^{k-1})^{-1} - N \\ & \leq (2\sqrt{c} + c) \sum_{i_0 \cdots i_m \in S} q(i_m | i_0^{m-1})^{-1}. \end{aligned} \quad (47)$$

Analogously, from (9) we can get

$$\liminf_{n \rightarrow \infty} \frac{1}{\sigma_n} \sum_{k=m}^{[\sigma_n]} q(X_k | X_{k-m}^{k-1})^{-1} - N \geq -2\sqrt{c} \sum_{i_0 \cdots i_m \in S} q(i_m | i_0^{m-1})^{-1}. \quad (48)$$

(45), (46) follow from (47) and (48), respectively. ◀

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